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For: METHOD TO EVALUATE WHETHER A TIME DELAY IS BETTER THAN A  
TIME LIMIT

Assistant Commissioner for Patents  
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**CLAIM OF PRIORITY UNDER 35 U.S.C. § 119(a)-(d)**

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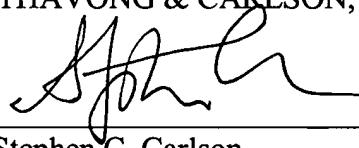
The benefit of any foreign patent application(s) listed below is hereby claimed under Title 35, United States Code, Section 119(a)-(d):

European Patent App. No. 02 023 410.0, filed October 18, 2002

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Respectfully Submitted,

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Die angehefteten Unterlagen stimmen mit der ursprünglich eingereichten Fassung der auf dem nächsten Blatt bezeichneten europäischen Patent anmeldung überein.

The attached documents are exact copies of the European patent application described on the following page, as originally filed.

Les documents fixés à cette attestation sont conformes à la version initialement déposée de la demande de brevet européen spécifiée à la page suivante.

**Patentanmeldung Nr.   Patent application No.   Demande de brevet n°**

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Der Präsident des Europäischen Patentamts;  
Im Auftrag

For the President of the European Patent Office  
Le Président de l'Office européen des brevets  
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Bezeichnung der Erfindung/Title of the invention/Titre de l'invention:  
(Falls die Bezeichnung der Erfindung nicht angegeben ist, siehe Beschreibung.  
If no title is shown please refer to the description.  
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Method to evaluate whether a time delay is better than a time limit

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Method to evaluate whether a time delay is better than a time limit

5

The invention concerns a method to evaluate whether a time delay as an implementation dependent parameter is better than a statistically defined soft time limit. The invention is especially applied to a class of measurements 10 measuring the delay time from a cell quality swap, generated by a system simulator, to the registration message, generated by the user equipment, a mobile station for example. In a mobile communication system the mobile station (user equipment) should make a cell reselection 15 or handover to another base station of another communication cell if the quality of communication with the current base station of the current cell (cell quality) decreases and the communication quality with an other base station of an other cell increases over the 20 quality of the current base station.

Such a soft handover handled by a mobile station in a communication system with several base stations for a mobile system of the third generation using Code Division 25 Multiple Access (CDMA) is known from US 5,267,261 for example.

The communication standard defines a maximum delay time from the swap cell quality until the time where the user 30 equipment issues a registration message in order to register to the other base station. However, this time delay TD is not defined as a hard limit, i. e. the user equipment would not fail the test if the delay time exceeds the time limit only a single time, but is defined 35 as a soft time limit, i. e. the user equipment shall fulfil the test requirement for a certain percentage (for example 90%) of the cases in repeated measurements. The pass fail decision of the test-equipment shall be done

with a certain quality, for example 3% wrong decision risk.

From the present state of the art it is not known how to  
5 deal with such statistically defined soft time limits for  
repeated tests.

Such a test is time consuming. This is illustrated by the  
following example: The success ratio SR in this example is  
10 90%. The time limit in this example is 8s. Due to  
statistical significance it is not enough to observe 9  
successes in 10 measurements. For statistical significance  
it is usual to observe  $200*9$  successes in  $200*10$   
measurements. This test lasts 4.5 hours.

15 It is the object of the present invention to provide an  
effective method to make a pass fail decision, especially  
on repeated time delay measurements, against a  
statistically defined soft time limit.

20 It is further the object of the present invention to  
provide an effective method to make the pass fail decision  
in the minimum possible test time preserving the  
statistical significance.

25 The object is solved alternatively by the features of  
claim 1, 2, 7 or 8.

30 The depended claims concern further developments of the  
invention. Claims 20 to 23 cover the respective digital  
storage mediums computer program and computer program  
product.

35 The invention is further described with respect to the  
drawings. In the drawings

Fig. 1 shows the scenario of cell swap and measurement  
of delay time;

Fig. 2 shows a block diagram of the measurement set up;

Fig. 3 shows a diagram to illustrate the inventive method,

5

Fig. 4 shows the referenced Error Ratio  $e_{norm}$  as a function of the measured errors  $n_e$ ,

Fig. 5 shows a diagram to illustrate a measurement using a first embodiment of the inventive method,

10 Fig. 6 shows a diagram to illustrate a measurement using a second embodiment of the inventive method and

15 Fig. 7 shows a diagram illustrating the position at the end of the test using the first embodiment of the inventive method as a function of probability.

20 Fig. 1 shows the scenario of the swap and measurement of time delay. Fig. 2 shows the respective measurement set up.

25

A system simulator SS simulates the two cells cell 1 and cell 2 of a cellular mobile communication system. A User Equipment UE (mobile station) is registered to one of the cells of the cellular mobile communication system. The

30 system simulator SS swaps the cell quality at specific cyclic time points T1 and T2. From time T1 until time T2 cell 1 has the better cell quality compared to cell 2. For example the field strength received by the User Equipment UE from cell 1 is 3dB higher than received from cell 2.

35 From T2 to T1 the cell quality of cell 2 is better as the cell quality of cell 1. The User Equipment UE should recognise the cell swap CS and should react to the cell swap CS by issuing a registration message RM. With the registration message RM the User Equipment UE registers to

the cell serving with better cell quality. The User Equipment UE needs a specific time delay TD to evaluate and recognise the cell swap CS and to react thereon with the registration message RM. This time delay TD is 5 measured for a specific number of tests by the system simulator SS and the result is transferred to the controller CO, as shown in Fig. 2. According to the inventive method the time delays TD are measured. Each time delay TD is hard decided, if it is worse or better 10 than the time limit. Then it is evaluated whether the delays are better than the time-limit in 90% of the cases. After sufficient repetitions an overall pass fail decision with a wrong decision risk of e.g.3% is done in a minimum possible test time.

15

The following realisation is exploited:

The occurrence of the bad delay in all delays, measured, is statistically independent, or can easily made statistically independent, by inserting sufficient 20 recovery time after a registration event. Statistically independent events are determined by the binomial distribution for any error ratio. (Error Ratio  $ER = 1-SR$  Success Ratio)

25 For small error ratios the binomial distribution is approximated by the Poisson distribution. The use of both distributions give meaningful results: The Poisson distribution give optimal results in terms of test time and statistical significance for small error ratios, and 30 the Poisson distribution is well supported in mathematical software tools.

For high error ratios the Poisson distribution give sub-optimal results. The results are conservative in the 35 sense: Test time is longer (worse) than optimal. Wrong decision risk is lower (better) than predefined.

The binomial distribution is optimal in terms of test time and statistical significance for any error ratio. However

the binomial distribution is not so well supported in standard mathematical software tools. Both derivations are described in the following.

- 5 According to the present invention, the worst possible likelihood distribution  $PD_{high}$  and/or the best possible likelihood distribution  $PD_{low}$  are obtained by the formulas given in claim 1 or 2 and 7 or 8, respectively. From these likelihood distributions, the average number  $NE_{high}$  or
- 10  $NE_{low}$  is obtained and compared with  $NE_{limit}$ , respectively. If  $NE_{limit}$  is higher than the average number  $NE_{high}$  or smaller than the average number  $NE_{low}$ , the test is stopped and it is decided that the device has early passed or early failed, respectively. During the test a single step
- 15 wrong decision probability for a preliminary ER stage can be used, which is smaller than the probability for the entire test.

- 20 The Poisson distribution preferably used for the present invention, is best adapted to a small number of "bad delay times". A "bad delay time" is defined as a delay time, exceeding the time limit. For example a measured time delay exceeding the time limit of 8s would be a "bad time delay", also referred to as an "error". A measured time
- 25 delay TD below the time limit, of for example 8s, would be a "good time delay".

- 30 In the following, a test procedure together with the early fail condition and the early pass condition is derived with mathematical methods.

- 35 The ER test procedure is an iterative process: The number ns of delays TD and the number of errors ne are accumulated from the beginning of the test to calculate a preliminary ER (Error Ratio). Time delays TD up to the next error are used together with the past time delays TD to calculate the next preliminary ER. Every preliminary ER is tested against an early pass or early fail bound. The formulas to derive an early pass or early fail condition

contain a single step wrong decision probability D. This is valid only for one specific preliminary ER stage. However, it is meaningful to have a wrong decision probability F for the entire test ( $F > D$ ). The application 5 proposes to use a single step wrong decision probability D, which is smaller than the probability for the entire test and proposes a method to derive D from a given F.

Due to the nature of the test, namely discrete error 10 events, the early stop conditions are declared not valid, when a fractional number of bad time delays  $< 1$  are used to calculate the early stop limits. The application contains proposals, how to conduct the test at this 15 undefined areas. The proposals are conservative (not risky). A DUT on the ER-limit does not achieve any early stop condition. The application proposes to stop the test unconditioned at a specific number K (for example 200) bad time delays. As this proposal contains different measurement qualities, an additional proposal is made to 20 overcome this drawback.

Based on a measurement at a preliminary ER stage pERs, a 25 confidence range CR around this measurement is derived. It has the property that with high probability the real result can be found in this range.

The confidence range CR is compared with the specified ER limit. From the result a diagram is derived containing the early fail and the early pass condition.

30 With a finite number of delays  $ns$ , the final Error Ratio ER cannot be determined exactly. Applying a finite number of delays  $ns$ , a number of bad time delays  $ne$  is measured.  $ne/ns = er$  is the preliminary Error Ratio.

35 In a test at a preliminary ER stage pERs a finite number  $ns$  of measured time delays TD is applied, and a number of bad time delays  $ne$  is measured.  $ne$  is connected with a certain differential probability in the Poisson

distribution. The probability and the position in the distribution conducting just one test at a preliminary ER stage pERs is not known.

5 Repeating this test infinite times, applying repeatedly the same ns, the complete Poisson distribution is obtained. The average number (mean value) of bad time delays is NE. NE/ns is the real ER at the DUT. The Poisson distribution has the variable ne and is characterised by

10 the parameter NE, the average or mean value. Real probabilities to find ne between two limits are calculated by integrating between such limits. The width of the Poisson distribution increases proportional to SQR(NE), that means, it increases absolutely, but decreases

15 relatively.

In a test at a preliminary ER stage pERs ns time delays TD are applied and ne bad time delays are measured. The result can be a member of different Poisson distributions

20 each characterized by another parameter NE. Two of them are given as follows:

The worst possible distribution  $NE_{high}$ , containing the measured ne with the probability  $D_1$ , is given by the

25 inverse cumulative operation

$$D_1 = \int_0^{ne} PD_{high}(ni, NE_{high}) dni \quad (1)$$

In the example  $D_1$  is  $0.002=0.2\%$   $PD_{high}$  is the wanted

30 Poisson distribution with the variable ni. ne is the measured number of bad time delays.

The best possible distributions  $NE_{low}$ , containing the measured ne with the probability  $D_2$  is given by the

35 inverse cumulative operation

$$D_2 = \int_{ne}^{\infty} PD_{low}(ni, NE_{low}) dni \quad (2)$$

In the example  $D_2$  is equal  $D_1$  and it is  $D = D_1 = D_2 = 0.002 = 0.2\%$ .

5

To illustrate the meaning of the range between  $NE_{low}$  and  $NE_{high}$  refer to Fig. 3. Fig. 3 shows the likelihood density  $PD$  as a function of the measured number of bad time delays  $ne$ . In the example, the actual detected number 10 of bad time delays  $ne$  within the measured number  $ns$  of time delays  $TD$  is 10. The likelihood distribution of the bad time delays is not known. The worst possible likelihood distribution  $PD_{high}$  under all possible likelihood distributions as well as the best possible 15 likelihood distribution  $PD_{low}$  under all possible likelihood distributions are shown. The worst possible likelihood distribution  $PD_{high}$  is characterized in that the integral from 0 to  $ne = 10$  gives a total probability of  $D_1 = 0.002$ . The best possible likelihood distribution 20  $PD_{low}$  is characterized in that the integral from  $ne = 10$  to  $\infty$  gives a total probability of  $D_2 = 0.002$ . In the preferred embodiment  $D_1$  is equal to  $D_2$ , i.e.  $D_1 = D_2 = 0.002 = 0.2\%$ . After having obtained the likelihood distribution  $PD_{high}$  and  $PD_{low}$  from formulas (1) and (2), 25 the average values or mean values  $NE_{high}$  for the likelihood distribution  $PD_{high}$  and  $NE_{low}$  for the likelihood distribution  $PD_{low}$  can be obtained. The range between the mean values  $NE_{low}$  and  $NE_{high}$  is the confidence range CR indicated in Fig. 3.

30

In the case the measured value  $ne$  is a rather untypical result (in the example just 0.2% probability) nevertheless the real result  $NE$  can still be found in this range, called confidence range CR.

35

The probabilities  $D_1$  and  $D_2$  in (1) and (2) can be independent, but preferable they are dependent and equal ( $D = D_1 = D_2$ ).

For the Poisson distribution  $NE_{low}$  and  $NE_{high}$  are the result of the inverse cumulative operations above and are shown in formulas (3) and (4) respectively. The inputs are

5 the number of bad time delays  $ne$ , measured in this test, and the probabilities  $D$  and  $C = 1 - D$ . The Output is  $NE$ , the parameter describing the average of the Poisson distribution.

10 The following example is illustrated in Fig. 3 ( $D = D_1 = D_2$ ):

$$NE_{low} = \frac{qchisq(D, 2 \cdot ne)}{2} \quad (3)$$

$$15 \quad NE_{high} = \frac{qchisq(C, 2 \cdot ne)}{2} \quad (4)$$

Example:

20 Number of bad time delays:  $ne=10$   
 Probability:  $D=0.002 \quad C=0.998$

Result:

25  $NE_{low} = 3.26$   
 $NE_{high} = 22.98$

Interpretation:

Having measured  $ne=10$  bad time delays in a test with preliminary ER stage pERs, then at a low probability

30  $D=0.002$  the average number of bad time delays  $NE$  in this test is outside the range from 3.26 to 22.98 or with a high probability  $C=0.998$  inside this range from 3.26 to 22.98.

35 Such as the width of the Poisson distribution, the confidence range  $CR$  increases proportional to  $SQR(ne)$ , that means, it increases absolutely, but decreases relatively.

If the entire confidence range CR, calculated from a single result ne, is found on the good side ( $NE_{limit} > NE_{high}$ ) of the specified  $NE_{limit}$  we can state: With high probability C, the final result NE is better than  $NE_{limit}$ . Whereby  $NE_{limit}$  is given by

$$NE_{limit} = ER_{limit} \cdot ns \quad (5)$$

10 and  $ER_{limit}$  is the Error Ratio allowable for the device and obtained by an ideal long test with an infinite high number ns of time delays TD.

15 If the entire confidence range CR, calculated from a single result ne, is found on the bad side ( $NE_{limit} < NE_{low}$ ) of the specified  $NE_{limit}$  we can state: With high probability C, the final result NE is worse than the  $ER_{limit}$ .

20 With each new time delay TD a new test is considered, reusing all former results. With each new test the preliminary data for ns, ne and er is updated. For each new test the confidence range CR is calculated and checked against  $NE_{limit}$ .

25 Once the entire confidence range CR is found on the good side of the specified limit ( $NE_{limit} > NE_{high}$ ), an early pass is allowed. Once the entire confidence range CR is found on the bad side of the specified limit ( $NE_{limit} < NE_{low}$ ) an early fail is allowed. If the confidence range CR is found on both sides of the specified limit ( $NE_{low} < NE_{limit} < NE_{high}$ ), it is evident neither to pass nor to fail the DUT early.

35 Fig. 3 illustrates the above conditions. Of course,  $NE_{limit}$  is a fixed value not altering during the test, but  $NE_{low}$  and  $NE_{high}$  as well as the confidence range CR are altering during the test. For reasons of illustration, however, the three possibilities of the possible positions of the

confidence range CR with respect to the constant  $NE_{limit}$  are drawn for the same example in Fig. 3.

5 The above can be described by the following formulas:

The current number  $ns$  of time delays TD is calculated from the preliminary Error Ratio  $er$  and the preliminary number of bad time delays  $ne$

10  $er = ne/ns$  (6)

The specified ER expressed with number of time delays  $ns$  and number of bad time delays  $NE$  is

15  $ER_{limit} = NE_{limit} / ns$  (7)

for abbreviation in the formula:

20  $er_{norm} = er/ER_{limit} = ne/ NE_{limit}$   
(normalised er) (8)

Early pass stipulates:

25  $NE_{high} < NE_{limit}$  (9)

Early fail stipulates:

$$NE_{low} > NE_{limit} \quad (10)$$

30 Formula for the early pass limit:

$$er_{norm} = \frac{ne}{NE_{high}} \quad (11)$$

35 This is the lower curve ( $ernorm_{pass} (ne, C)$ ) in Fig. 4, which shows  $er_{norm}$  as a function of  $ne$ .

Formula for the early fail limit:

$$er_{norm} = \frac{ne}{NE_{low}} \quad (12)$$

This is the upper curve ( $er_{norm, fail}$  (ne, D)) in Fig. 4.

5

As the early pass limit is not defined for  $ne = 0$  (normally the case at the very beginning of the test for a good DUT), an artificial bad time delay event with the first measured time delay TD can be introduced. When the 10 first real bad time delay event occurs, the artificial bad time delay is replaced by this real one. This gives the shortest possible measurement time for an ideal good DUT. For example  $ns=50$  for  $ER_{limit} = 0.1$  and probability  $D = D_1 = D_2 = 0.2\%$ . (6 minutes 20s)

15

As the early fail limit uses  $NE_{low} < 1$  for small  $ne < k$  (in the example below  $k=5$ ) due to a decision problem at a fractional error, the early fail limit at  $ne=k$  is extended with a vertical line upwards. This ensures that a very bad 20 DUT hits the early fail limit in any case after a few delays, approx. 10 in the example (1 minute 20s). In other words, the test is not stopped as long as  $ne$  is smaller than  $k$ .

25 With each new time delay TD a new test is considered, reusing all former results. With each new test the preliminary data for  $ns$ ,  $ne$  and  $er$  and  $er_{norm}$  are updated and a  $er_{norm}/ne$  coordinate is entered into the  $er_{norm}$ -diagram. This is shown in Fig. 5. Once the trajectory 30 crosses the early fail limit ( $er_{norm}$  (ne, D)) or the early pass limit ( $er_{norm}$  (ne, C)) the test may be stopped and the conclusion of early fail or early pass may be drawn based on this instant.

35 Fig. 5 shows the curves for early fail and early pass.  $er_{norm}$  is shown as a function of the number of bad time delays  $ne$ . For the simple example demonstrated in Fig. 5, it is  $ER_{limit} = 0.2 = 1/5$  and the final Error Ratio  $ER =$

0.25 (1/4). The test starts with the first time delay TD, for which no error (bad time delay) is detected. For the second time delay TD, a first error (bad time delay) is detected and the preliminary Error Ratio  $er = ne/ns = 1/2$

5 and  $er_{norm} = er/ER_{limit}$  becomes  $1/2 : 1/5 = 5/2$ .  $er_{norm}$  after the second time delay TD is marked with a cross a in Fig. 5. For the third, fourth and fifth time delay TD, no further error (bad time delay) occurs and  $er_{norm}$  subsequently becomes  $5/3$ ,  $5/4$  and  $5/5$ , respectively, which

10 is marked with the crosses b, c and d in Fig. 5, respectively. The sixth time delay TD brings a new error (bad time delay) and ne becomes 2. Consequently,  $er = ne/ns$  becomes  $2/6$  and  $er_{norm}$  becomes  $10/6$ . This situation is marked with cross e in Fig. 5. For the seventh, eighth

15 and ninth delay TD, no further error (bad time delay) occurs and the situation after the seventh, eighth and ninth delay TD is marked with crosses f, g, h in Fig. 5, respectively. The tenth time delay TD brings a third error (bad time delay). Consequently, er becomes  $3/10$  and  $er_{norm}$

20 becomes  $15/10$ . This situation is marked with cross i in Fig. 5. As can be seen from Fig. 5, the trajectory is between the early fail curve and the early pass curve at the beginning of the test, but converges to a line Z, which crosses the early fail curve after about forty

25 errors. After forty errors, it can thus be decided that the tested DUT early fails the test.

The calculations above give optimal results in terms of test-time for small error ratios /high success ratios. If

30 this is not the case, the approach with the Poisson distribution can be used nevertheless. For any error ratio/success ratio the binomial distribution is valid. The Poisson distribution is wider than the equivalent binomial distribution, hence the confidence range is

35 wider, hence the space between the early pass and the early fail limit is wider, hence the result is not optimal in terms of test-time, but conservative is the following sense: the test-time is longer than optimal and the wrong decision risk is lower than predefined.

To get optimal results also for higher error ratios/lower success ratios the binomial distribution must be applied as follows:

5 The binomial distribution has the following form:  
 $dbinom(ne, ns, ER)$   
 with  
 ne: number of errors (bad time delays) detected,  
 ns, number of time delays TD, measured,  
 10 (ne/ns = preliminary error ratio er)  
 ER: the mean value of the distribution is the error ratio

15 The worst possible distribution with respect to a measured er=ne/ns is defined by the inverse cumulative operation:

$$D = \int_0^{ne} dbinom(ni, ns, ER_{high}) dni \quad (13)$$

with  
 D: wrong decision risk  
 20 ni: integration variable  
 ne: number of errors (bad time delays) currently detected,  
 simultaneously integration limit  
 ns: number of time delays TD, currently measured  
 ER<sub>high</sub>: The parameter to be tuned, to make the integral  
 25 consistent. This is the wanted value.  
 The result has the form:  $ER_{high} = qbinom(C, ne, ns)$   
 with  $C=1-D$ , the complementary wrong decision risk is equal  
 to the confidence level.

30 The best possible distribution is defined by the inverse cumulative operation:

$$D = \int_{ne}^n dbinom(ni, ns, ER_{low}) dni \quad (14)$$

35 with

ne: number of errors (bad time delays) currently detected, simultaneously integration limit

ns: number of time delays TD, currently measured, simultaneously integration limit

5 ER<sub>low</sub>: The parameter to be tuned, to make the integral consistent. This is the wanted value.

The result has the form:  $ER_{low} = qbinom(D, ne, ns)$

10 Another inverse cumulative inversion, than the one above, is frequently offered with standard mathematical software tools. The form is similar:  $qbinom(D, \dots)$  (at least one other parameters than above).

15 The confidence range is the range between ER<sub>high</sub> and ER<sub>low</sub>. The normalized er<sub>norm</sub> is the measured (preliminary) er=ne/ns divided by ER limit

Early pass stipulates:

$$ER_{high} < ER_{limit}$$

20

Early fail stipulates:

$$ER_{low} > ER_{limit}$$

25 In analogy to the Poisson approach, the early fail limit is

$$er_{norm\ fail} = er/ER_{low} \quad (15)$$

The early pass limit is

30

$$er_{norm\ pass} = er/ER_{high} \quad (16)$$

The following consideration is valid for both, the Poisson and the binomial approach:

35 If no early stop occurs the ER test may be stopped, after the following condition is valid:

$$ne \geq K \quad (17)$$

and the DUT shall be passed, if  $ns$  is sufficiently high.  $K$  is a maximum number of errors (bad time delays). For example  $K$  can be 200.

5 If the trajectory neither crosses the early fail curve nor the early pass curve after  $K$  (for example 200) errors (bad time delays) have occurred, the DUT can be finally passed. If the DUT, however, is rather good or rather bad, the tests can be stopped much earlier, long before the  $K = 200$

10 errors (bad time delays) have occurred. This significantly shortens the total test time.

In the above embodiment early fail means: a DUT, which survived to certain preliminary ER stage, is failed and a probability of 0.2 % that it is actually better than the ER-limit is accepted. Further early pass means: the DUT, which survived to certain preliminary ER stage, is passed and a probability of 0.2 % that it is actually worse than the ER-limit is accepted. If the test is stopped at 200

20 errors (bad time delays) the DUT is passed without any early fail or early pass condition arbitrarily. It can cross the vertical 200 error line in Fig. 4 at different heights.

25 Each height is connected with a certain statistical interpretation: The probability to have a DUT better (worse) than the ER-limit is indicated in Fig. 7. The vertical in Fig. 7 shows the position in Fig. 4 at the end of the test. The horizontal line in Fig. 7 shows the 30 respective probability. Hence such a test has different qualities, a better one at the early pass fail limits, and a worse and variable one at the unconditioned stop at 200 errors

35 The following embodiment overcomes this drawback and additionally accelerates the test. This is done by a meaningful redefinition of the early pass limit maintaining the early fail limit. Early pass means now: A DUT, which survived to certain preliminary ER stage, is

passed and a probability of 0.2 % that it is actually worse than  $M$  times the specified ER-limit ( $M > 1$ ) is accepted. This is a bad DUT ER-limit. This shifts the early pass limit upwards in Fig. 4 as shown in Fig. 6. 5  $er_{norm} \text{ pass } (ne, C)$  in Fig. 4 becomes  $er_{normbad} \text{ pass } (ne, C)$  in Fig. 6.  $er_{norm} \text{ fail } (ne, D)$  remains unchanged. Now it is

$$NE_{\text{limit}, M} = ER_{\text{limit}} \cdot M \cdot ns \quad \text{for Poisson}$$

$$ER_{\text{limit}, M} = ER_{\text{limit}} \cdot M \quad \text{for Binomial}$$

10 and an early pass is allowed, if

$$NE_{\text{limit}} \cdot M = NE_{\text{limit}, M} > NE_{\text{high}}. \quad (\text{Poisson})$$

$$ER_{\text{limit}} \cdot M = ER_{\text{limit}, M} > ER_{\text{high}} \quad (\text{binomial}) \quad (18)$$

There are three high level input parameters for the test:

15 - Specified error limit e.g. 10%

- Probability to make a wrong decision in the entire test (proposal:  $F = 3 \%$ )

- Definition of a bad DUT ( $ER_{\text{limit}} \cdot M$ )

20 From this the working parameters are derived:

D, the wrong decision risk for a single test step, is derived from  $F$ , the wrong decision risk for the entire test. The method to do this is described below.

25 With D the early pass and fail limits can be calculated according to formula 11, 12, 15, 16.

30 The intersection co-ordinates of the early pass and the early fail limit are the maximum number of errors (bad time delay) and the error ratio test limit TL (specified error limit  $< TL < M * \text{specified error limit}$ ).

There are infinite possibilities to overcome the above mentioned drawback with different test qualities.

35

In the example above the open end between the pass fail-limits was originally declared pass, and time was saved by multiplying the early pass limit with  $M$  ( $M > 1$ ), shifting it upwards such that the early fail and the early pass curve

intersect at 200 errors (200 bad time delays: example from above). Thus only a DUT, bad with high probability, is failed (customer risk).

5 The complementary method is: The open end between the early pass fail limits is declared fail, and time is saved by multiplying the early fail limit with  $m$  ( $0 < m < 1$ ), shifting it downwards, such that the early fail and the early pass curve intersect at 200 errors (200 bad time

10 delays: example from above). Thus only a DUT, good with high probability, is passed (manufacturer risk).

The compromise method is: The open end between the early pass fail limits is partitioned in any ratio: the upper

15 part is declared fail and the lower part is declared pass. Time is saved by multiplying the early fail limit with  $m$  ( $0 < m < 1$ ) and such shifting it downwards and by multiplying the early pass limit with  $M$  ( $M > 1$ ) and such shifting it upwards. So the early fail and the early pass curve

20 intersect at 200 errors (200 bad time delays: example from above).

With given  $D_1$  and  $D_2$  the early fail curve and the early pass curves in Fig. 5 and Fig. 4 or Fig. 6 can be

25 calculated before the test is started. During the test only  $er_{norm} = ne/NE_{limit}$  (for binomial:  $er_{norm} = er/ER_{limit}$ ) has to be calculated and to be compared with the early pass limit and the early fail limit as explained with respect to Fig. 5 and Fig. 6. Thus, no intensive

30 calculation has to be done during the test.

$D_1$  and  $D_2$  above describe the single step wrong decision probability for a DUT, which survived without early stop to a certain preliminary ER stage. For a real test it is

35 desirable to define in advance a wrong decision probability  $F$  for the entire test. For every step a fraction of DUTs leave the statistical totality due to single step wrong decision probability  $D$ . This accumulates to an amount  $F > D$ .

It is proposed to derive D from F by the following method:

Based on statistical independence, a large ensemble of  
 5 DUTs with ER on the ER-limit (limit-DUT) and with M\*ER  
 (bad DUT) is simulated and the entire test for the limit-  
 DUT ensemble and the bad DUT ensemble is run by simulation  
 against the early pass and early fail bound, with a free  
 10 D-parameter ( $D_1$  and  $D_2$ ). The simulation will show, that a  
 certain fraction F ( $D < F < 1$ ) of the large ensemble falsely  
 fails (limit-DUT) or falsely passes (bad DUT). This  
 represents the wrong decision probability F for the entire  
 test. D is tuned such that F corresponds to the predefined  
 wrong decision probability.

15

The initially mentioned wrong decision risk 3%  
 corresponds, as a result of the simulation, the single  
 step wrong decision risk of 0.2% applied in the formulas  
 and figures.

20

In retrospect the distributions are not any more exact  
 applying this simulation: Once a DUT leaves the  
 statistical totality due to early fail or early pass, the  
 subsequent distribution is changed. In so far the  
 25 initially selected distribution has its intended validity  
 only at the beginning of the test. Due to early passes and  
 early fails the distribution is truncated more and more  
 towards the end of the test. These truncated distributions  
 cannot be handled anymore analytically, and so it is  
 30 impossible to derive D from F analytically. It is possible  
 to derive each subsequent distribution by numerical  
 methods. Then it is possible to derive D from F  
 numerically. In the above proposed process, just a unique,  
 the initial distribution is applied as approximation  
 35 instead of a set of correct distributions. With a set of  
 correct distributions it is possible to calculate D from F  
 by:

$$F = 1 - (1 - D)^n$$

(19)

Generally it can be assumed:

$$F > D \geq 1 - (1-F)^{1/m} \quad (20)$$

5

This formula can serve as an infimum for D, if the method is based on simulation using the initial distribution.

5 1. Method for testing the time delay error ratio ER of a device against a maximal allowable time delay error ratio ER<sub>limit</sub> with an early pass criterion, whereby the early pass criterion is allowed to be wrong only by a small probability D<sub>1</sub>, with the following steps

10 - measuring n<sub>s</sub> time delays (TD) of the device, thereby detecting n<sub>e</sub> bad time delays, which exceed a certain time limit, of these n<sub>s</sub> time delays (TD),  
 - assuming that the likelihood distribution giving the distribution of the number n<sub>i</sub> of bad time delays in a

15 fixed number of samples of time delays (TD) is PD(n<sub>i</sub>, NE), wherein NE is the average number of bad time delays, obtaining PD<sub>high</sub> from

$$D_1 = \int_0^{n_e} PD_{high}(n_i, NE_{high}) d n_i$$

20 wherein PD<sub>high</sub> is the worst possible likelihood distribution containing the measured n<sub>e</sub> bad time delays with the probability D<sub>1</sub>,  
 - obtaining the average number NE<sub>high</sub> of bad time delays  
 25 for the worst possible likelihood distribution PD<sub>high</sub>,  
 - comparing NE<sub>high</sub> with NE<sub>limit</sub> = ER<sub>limit</sub> · n<sub>s</sub>,  
 - if NE<sub>limit</sub> is higher than NE<sub>high</sub> stopping the test and deciding that the device has early passed the test and  
 - if NE<sub>limit</sub> is smaller than NE<sub>high</sub> continuing the test  
 30 whereby increasing n<sub>s</sub>.

2. Method for testing the time delay error ratio ER of a device against a maximal allowable time delay error ratio ER<sub>limit</sub> with an early pass criterion, whereby the early  
 35 pass criterion is allowed to be wrong only by a small probability F<sub>1</sub> for the entire test, with the following steps  
 - measuring n<sub>s</sub> time delays (TD) of the device, thereby

detecting  $n_e$  bad time delays, which exceed a certain time limit, of these  $n_s$  time delays (TD)

- assuming that the likelihood distribution giving the distribution of the number  $n_i$  of bad time delays in a
- 5 fixed number of samples of time delays (TD) is  $PD(n_i, NE)$ , wherein  $NE$  is the average number of bad time delays, obtaining  $PD_{high}$  from

$$D_1 = \int_0^{n_e} PD_{high}(n_i, NE_{high}) dn_i$$

- 10 wherein  $PD_{high}$  is the worst possible likelihood distribution containing the measured  $n_e$  bad time delays with a single step wrong decision probability  $D_1$  for a preliminary error ratio ER stage, whereby using a single step wrong decision probability  $D_1$  smaller than the
- 15 probability  $F_1$  for the entire test,

- obtaining the average number of  $NE_{high}$  of bad time delays for the worst possible likelihood distribution  $PD_{high}$ ,
- comparing  $NE_{high}$  with  $NE_{limit} = ER_{limit} \cdot n_s$ ,
- 20 - if  $NE_{limit}$  is higher than  $NE_{high}$  stopping the test and deciding that the device has early passed the test and
- if  $NE_{limit}$  is smaller than  $NE_{high}$  continuing the test whereby increasing  $n_s$ .

- 25 3. Method according to claim 1,  
characterized in that

the single step wrong decision probability  $D_1$  is in the range of

30  $F_1 > D_1 \geq 1 - (1 - F_1)^{1/n_e}$ .

- 4. Method according to any of claims 1 to 3,  
characterized in that
- 35 the likelihood distribution  $PD_{high}(n_i, NE)$  is the Poisson distribution.

- 5. Method according to any of claims 1 to 3,  
characterized in that

the likelihood distribution  $PD_{high}(ni, NE)$  is the binomial distribution.

6. Method according to any of claims 1 to 5,  
 5 characterized in that  
 for avoiding an undefined situation for  $ne=0$  starting the test with an artificial bad time delay  $ne=1$ , not incrementing  $ne$  then a first error occurs.

10 7. Method for testing the time delay error ratio  $ER$  of a device against a maximal allowable time delay error ratio  $ER_{limit}$  with an early fail criterion, whereby the early fail criterion is allowed to be wrong only by a small probability  $D_2$ , with the following steps  
 15 - measuring  $ns$  time delays (TD) of the device, thereby detecting  $ne$  bad time delays, which exceed a certain time limit, of these  $ns$  time delays (TD),  
 - assuming that the likelihood distribution giving the distribution of the number  $ni$  of bad time delays in a  
 20 fixed number of samples of time delays (TD) is  $PD(ni, NE)$ , wherein  $NE$  is the average number of bad time delays, obtaining  $PD_{low}$  from the

$$D_2 = \int_{ne}^{\infty} PD_{low}(ni, NE_{low}) dni$$

25 wherein  $PD_{low}$  is the best possible likelihood distribution containing the measured  $ne$  bad time delays with the probability  $D_2$ ,  
 - obtaining the average number  $NE_{low}$  bad time delays for  
 30 the best possible likelihood distribution  $PD_{low}$ ,  
 - comparing  $NE_{low}$  with  $NE_{limit} = ER_{limit} \cdot ns$ ,  
 - if  $NE_{limit}$  is smaller than  $NE_{low}$  stopping the test and deciding that the device has early failed the test and  
 - if  $NE_{limit}$  is higher than  $NE_{low}$  continuing the test  
 35 whereby increasing  $ns$ .

8. Method for testing the time delay error ratio  $ER$  of a device against a maximal allowable time delay error ratio

ER<sub>limit</sub> with an early fail criterion, whereby the early fail criterion is allowed to be wrong only by a small probability  $F_2$  for the entire test, with the following steps

5 - measuring  $ns$  time delays (TD) of the device, thereby detecting  $ne$  bad time delays, which exceed a certain time limit, of these  $ns$  time delays (TD),  
 - assuming that the likelihood distribution giving the distribution of the number  $ni$  of bad time delays in a  
 10 fixed number of samples of time delays (TD) is  $PD(ni, NE)$ , wherein  $NE$  is the average number of bad time delays, obtaining  $PD_{low}$  from

$$D_2 = \int_{ne}^{\infty} PD_{low}(ni, NE_{low}) dni$$

15 wherein  $PD_{low}$  is the best possible likelihood distribution containing the measured  $ne$  bad time delays with a single step wrong decision probability  $D_2$  for a preliminary error ratio ER stage, whereby using a single step wrong decision probability  $D_2$  smaller than the probability  $F_2$  for the  
 20 entire test,  
 - obtaining the average number  $NE_{low}$  bad time delays for the best possible likelihood distribution  $PD_{low}$ ,  
 - comparing  $NE_{low}$  with  $NE_{limit} = ER_{limit} \cdot ns$ ,  
 - if  $NE_{limit}$  is smaller than  $NE_{low}$  stopping the test and  
 25 deciding that the device has early failed the test and  
 - if  $NE_{limit}$  is higher than  $NE_{low}$  continuing the test whereby increasing  $ns$ .

9. Method according to claim 8,  
 30 characterized in that  
 the single step wrong decision probability  $D_2$  is in the range of

$$F_2 > D_2 \geq 1 - (1 - F_2)^{1/ne}.$$

35

10. Method according to any of claims 7 to 9,  
 characterized in that  
 the likelihood distribution  $PD_{low}(ni, NE)$  is the Poisson

distribution.

11. Method according to any of claims 7 to 9,  
characterized in that

5 the likelihood distribution  $PD_{low}(ni, NE)$  is the binomial distribution.

12. Method according to any of claims 7 to 11,  
characterized in that

10 for avoiding a undefined situation for neck, wherein k is a small number of bad time delays, not stopping the test as long as ne is smaller than k.

13. Method according to any of claims 7 to 12,  
15 characterized by

an additional early pass criterion, whereby the early pass criterion is allowed to be wrong only by a small probability  $D_1$ , with the following additional steps

20 - assuming that the likelihood distribution giving the distribution of the number of bad time delays  $ni$  in a fixed number of samples of time delays (TD) is  $PD(ni, NE)$ , wherein  $NE$  is the average number of bad time delays, obtaining  $PD_{high}$  from

$$25 \quad D_1 = \int_0^{ne} PD_{high}(ni, NE_{high}) dni$$

wherein  $PD_{high}$  is the worst possible likelihood distribution containing the measured  $ne$  bad time delays with the probability  $D_1$ ,

30 - obtaining the average number  $NE_{high}$  of bad time delays for the worst possible likelihood distribution  $PD_{high}$ ,  
- comparing  $NE_{high}$  with  $NE_{limit} = ER_{limit} \cdot ns$ ,  
- if  $NE_{limit}$  is higher than  $NE_{high}$  stopping the test and deciding that the device has early passed the test and  
35 - if  $NE_{limit}$  is smaller than  $NE_{high}$  continuing the test, whereby increasing  $ns$ .

14. Method according to any of claims 7 to 12,

**characterized by**

an additional early pass criterion, whereby the early pass criterion is allowed to be wrong only by a small probability  $D_1$ , with the following additional steps

5 - assuming that the likelihood distribution giving the distribution of the number of bad time delays  $n_i$  in a fixed number of samples of time delays (TD) is  $PD(n_i, NE)$ , wherein  $NE$  is the average number of bad time delays, obtaining  $PD_{high}$  from

10

$$D_1 = \int_0^{ne} PD_{high}(n_i, NE_{high}) dn_i$$

wherein  $PD_{high}$  is the worst possible likelihood distribution containing the measured  $ne$  bad time delays with the probability  $D_1$ ,

15 - obtaining the average number  $NE_{high}$  of bad time delays for the worst possible likelihood distribution  $PD_{high}$ ,  
 - comparing  $NE_{high}$  with  $NE_{limit,M} = ER_{limit} \cdot M \cdot ns$ , with  $M > 1$ ,  
 20 - if  $NE_{limit,M}$  is higher than  $NE_{high}$  stopping the test and deciding that the device has early passed the test and  
 - if  $NE_{limit,M}$  is smaller than  $NE_{high}$  continuing the test, whereby increasing  $ns$ .

25 15. Method according to claim 13 or 14,

**characterized in that**

the probability  $D_1$  for the wrong early pass criterion and the probability  $D_2$  for the wrong early fail criterion are equal ( $D_1=D_2$ ).

30

16. Method according to any of claims 7 to 12,

**characterized by**

an additional early pass criterion, whereby the early pass criterion is allowed to be wrong only by a small probability  $F_1$  for the entire test, with the following additional steps

35 - assuming that the likelihood distribution giving the distribution of the number  $n_i$  of bad time delays in a

fixed number of samples of time delays (TD) is  $PD(ni, NE)$ , wherein NE is the average number of bad time delays, obtaining  $PD_{high}$  from

$$5 \quad D_1 = \int_0^{ne} PD_{high}(ni, NE_{high}) dni$$

wherein  $PD_{high}$  is the worst possible likelihood distribution containing the measured ne bad time delays with the single step wrong decision probability  $D_1$  for a 10 preliminary error ratio ER stage, whereby using a single step wrong decision probability  $D_1$  smaller than the probability  $F_1$  for the entire test,

- obtaining the average number of  $NE_{high}$  bad time delays for the worst possible likelihood distribution  $PD_{high}$ ,
- 15 - comparing  $NE_{high}$  with  $NE_{limit} = ER_{limit} \cdot ns$ ,
- if  $NE_{limit}$  is higher than  $NE_{high}$  stopping the test and deciding that the device has early passed the test and
- if  $NE_{limit}$  is smaller than  $NE_{high}$  continuing the test, whereby increasing ns.

20

17. Method according to any of claims 7 to 12, characterized by

an additional early pass criterion, whereby the early pass criterion is allowed to be wrong only by a small 25 probability  $F_1$  for the entire test, with the following additional steps

- assuming that the likelihood distribution giving the distribution of the number  $ni$  of bad time delays in a fixed number of samples of time delays (TD) is  $PD(ni, NE)$ ,
- 30 wherein NE is the average number of bad time limits, obtaining  $PD_{high}$  from

$$D_1 = \int_0^{ne} PD_{high}(ni, NE_{high}) dni$$

35 wherein  $PD_{high}$  is the worst possible likelihood distribution containing the measured ne bad time delays with the single step wrong decision probability  $D_1$  for a

preliminary error ratio ER stage, whereby using a single step wrong decision probability  $D_1$  smaller than the probability  $F_1$  for the entire test,

- obtaining the average number  $NE_{high}$  of bad time delays
- 5 for the worst possible likelihood distribution  $PD_{high}$ ,
- comparing  $NE_{high}$  with  $NE_{limit,M} = ER_{limit} \cdot M \cdot ns$ , with  $M > 1$
- if  $NE_{limit,M}$  is higher than  $NE_{high}$  stopping the test and deciding that the device has early passed the test and
- 10 - if  $NE_{limit,M}$  is smaller than  $NE_{high}$  continuing the test, whereby increasing  $ns$ .

18. Method according to claim 16 or 17,  
characterized in that

15 the probability  $F_1$  for the wrong early pass criterion and the probability  $F_2$  for the wrong early fail criterion are equal ( $F_1=F_2$ ).

19. Method according to any of claims 7 to 18,  
20 characterized in that  
for avoiding a undefined situation for  $ne=0$  starting the test with an artificial bad time delay  $ne=1$  not incrementing  $ne$  then a first error occurs.

25 20. Digital storage medium with control signals electronically readable from the digital storage medium, which interact with a programmable computer or digital signal processor in a manner that all steps of the method according to any of claims 1 to 19 can be performed.

30 35 21. Computer-program-product with program-code-means stored on a machinereadable data carrier to perform all steps of any of claims 1 to 19, when the program is performed on a programmable computer or a digital signal processor.

22. Computer program with program-code-means to perform all steps of any of claims 1 to 19, when the program is performed on a programmable computer or a digital signal

processor.

23. Computer program with program-code-means to perform  
all steps of any of claims 1 to 19 when the program is  
5 stored on a machinereadable data carrier.

18. Okt. 2002

## Abstract

5 A method for testing the time delay error ratio  $ER$  of a device against a maximal allowable time delay error ratio  $ER_{limit}$  with an early pass and/or early fail criterion, whereby the early pass and/or early fail criterion is allowed to be wrong only by a small probability  $D$ . ns time  
10 delays  $TD$  of the device are measured, thereby  $ne$  bad time delays of these ns time delays  $TD$  are detected.  $PD_{high}$  and/or  $PD_{low}$  are obtained, whereby  $PD_{high}$  is the worst possible likelihood distribution and  $PD_{low}$  is the best possible likelihood distribution containing the measured  
15  $ne$  bad time delays with the probability  $D$ . The average numbers of erroneous bits  $NE_{high}$  and  $NE_{low}$  for  $PD_{high}$  and  $PD_{low}$  are obtained.  $NE_{high}$  and  $NE_{low}$  are compared with  $NE_{limit} = ER_{limit} \cdot ns$ . If  $NE_{limit}$  is higher than  $NE_{high}$  or  $NE_{limit}$  is lower than  $NE_{low}$  the test is stopped.

20

(Fig. 3)

Fig. 1

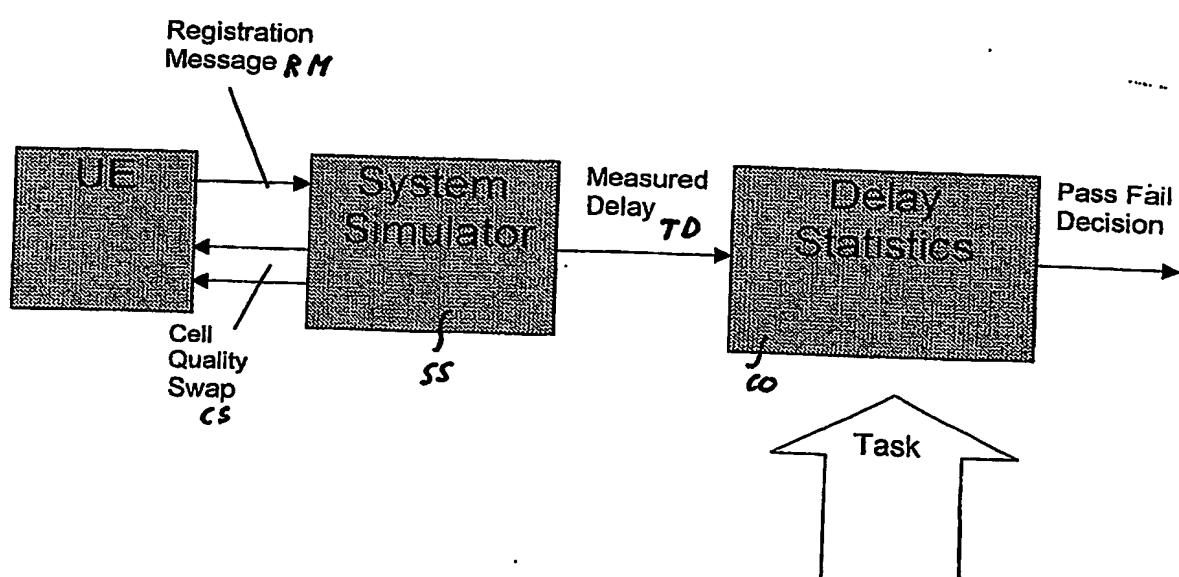
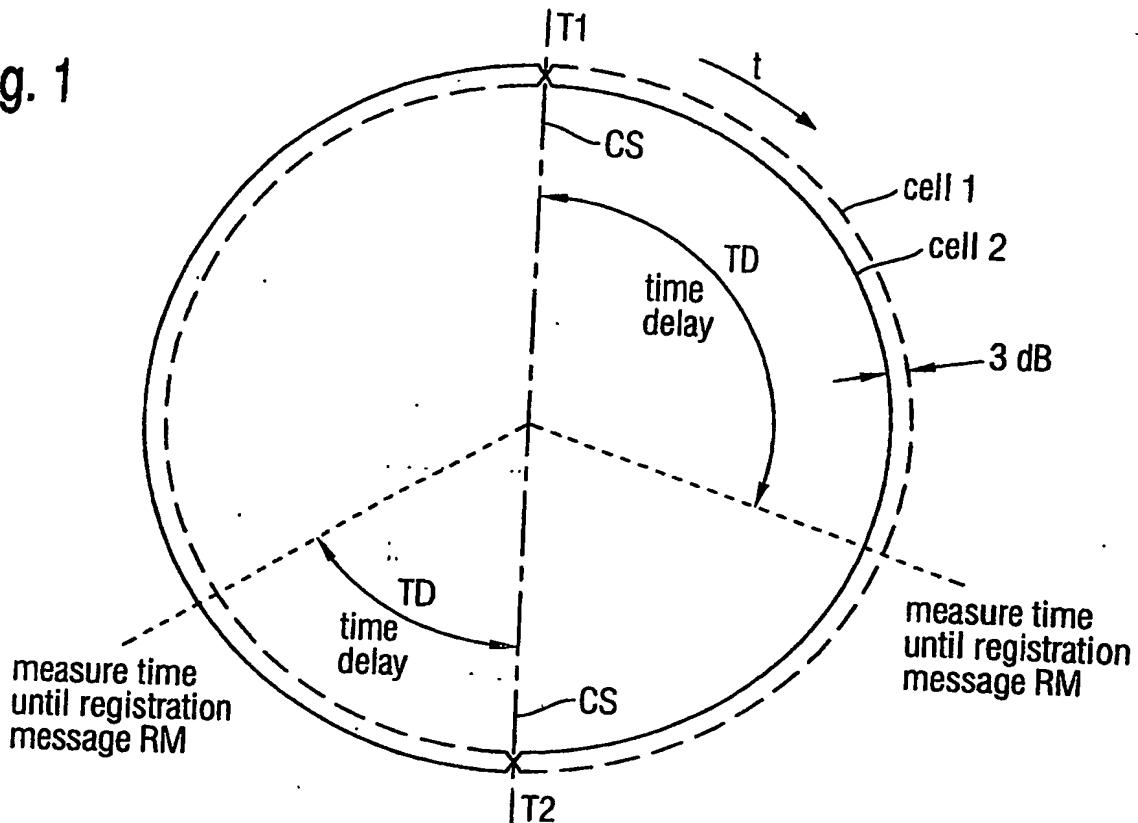


Fig. 2

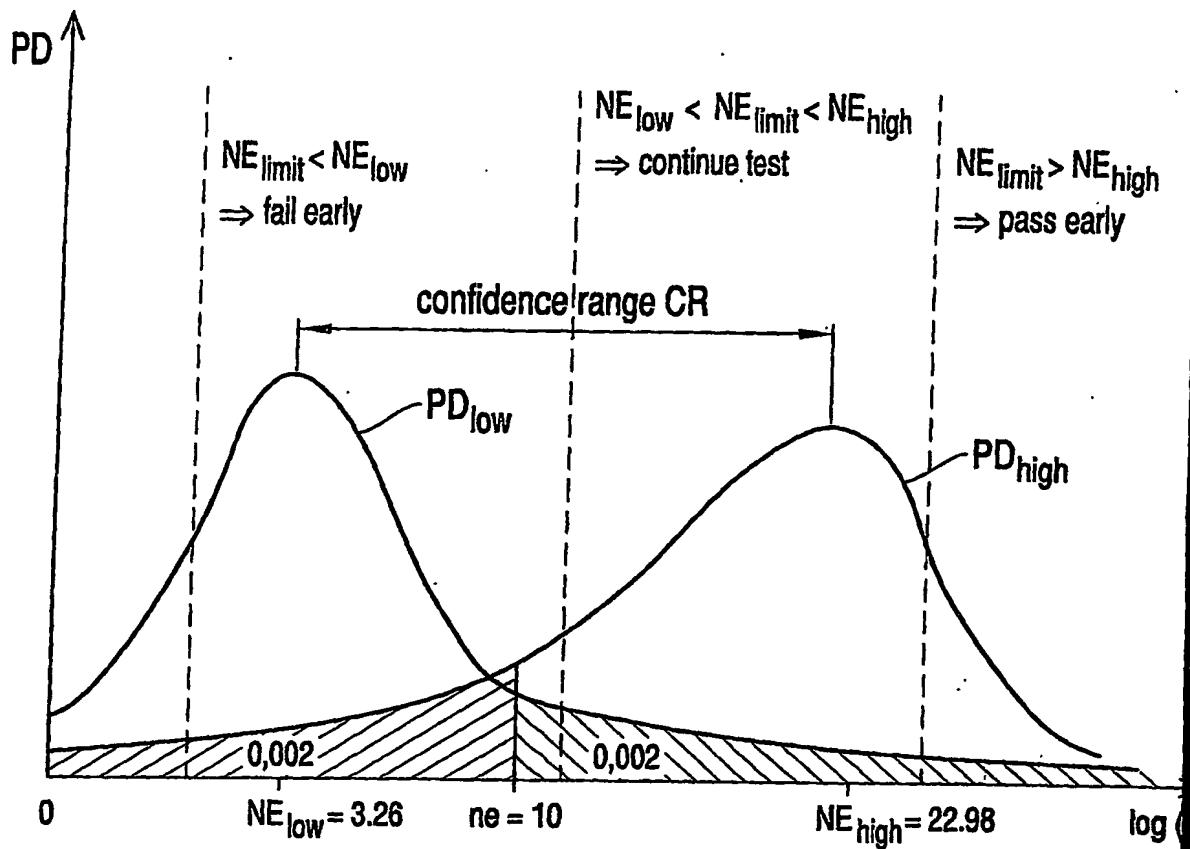


Fig. 3

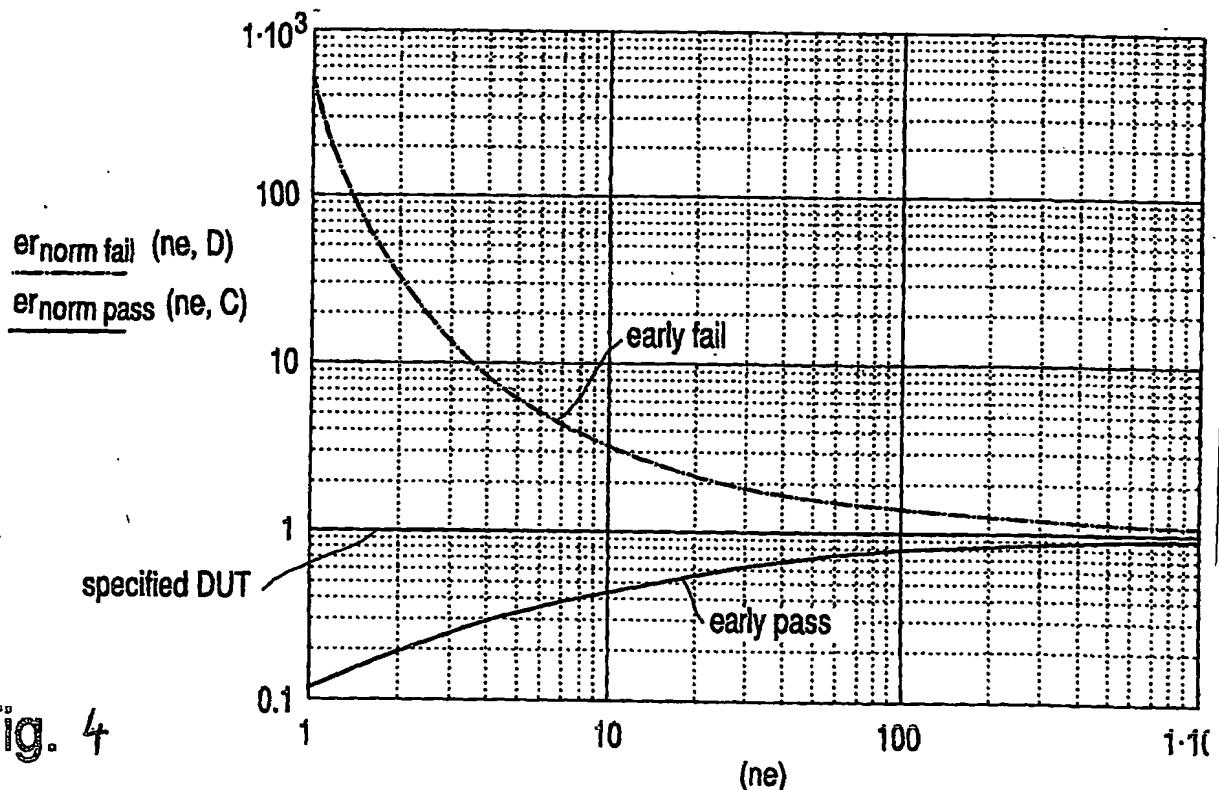


Fig. 4

3/4

ER limit = 0.2 (1/5) Actual ER = 0.25 (1/4)

Number of sample,	1	2	3	4	5	6	7	8	9	10 etc...
Sequence, 0: correct, 1: error	0	1	0	0	0	1	0	0	0	1
ER		1/2	1/3	1/4	1/5	2/6	2/7	2/8	2/9	3/10
ER norm		5/2	5/3	5/4	5/5	10/6	10/7	10/8	10/9	15/10

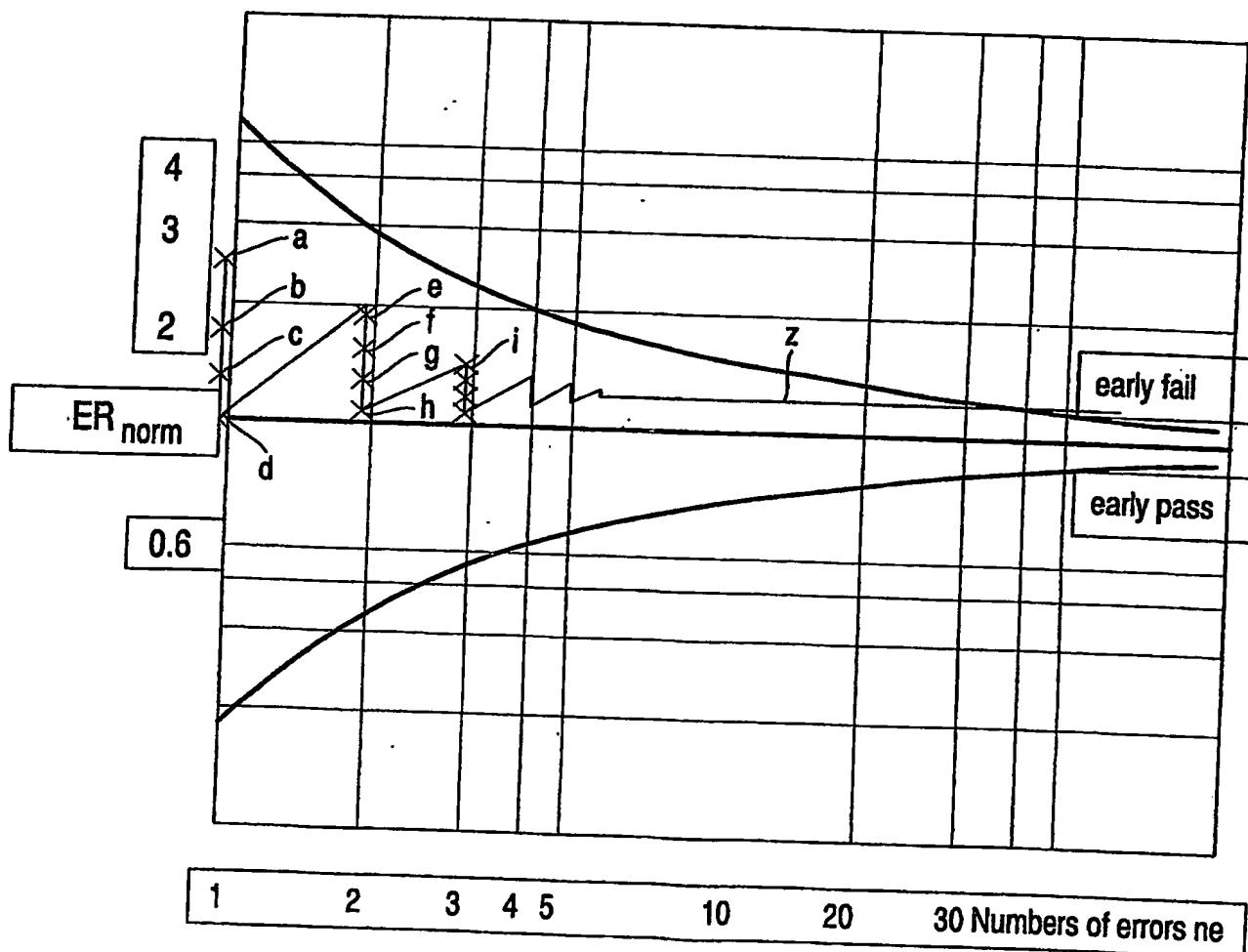


Fig. 5

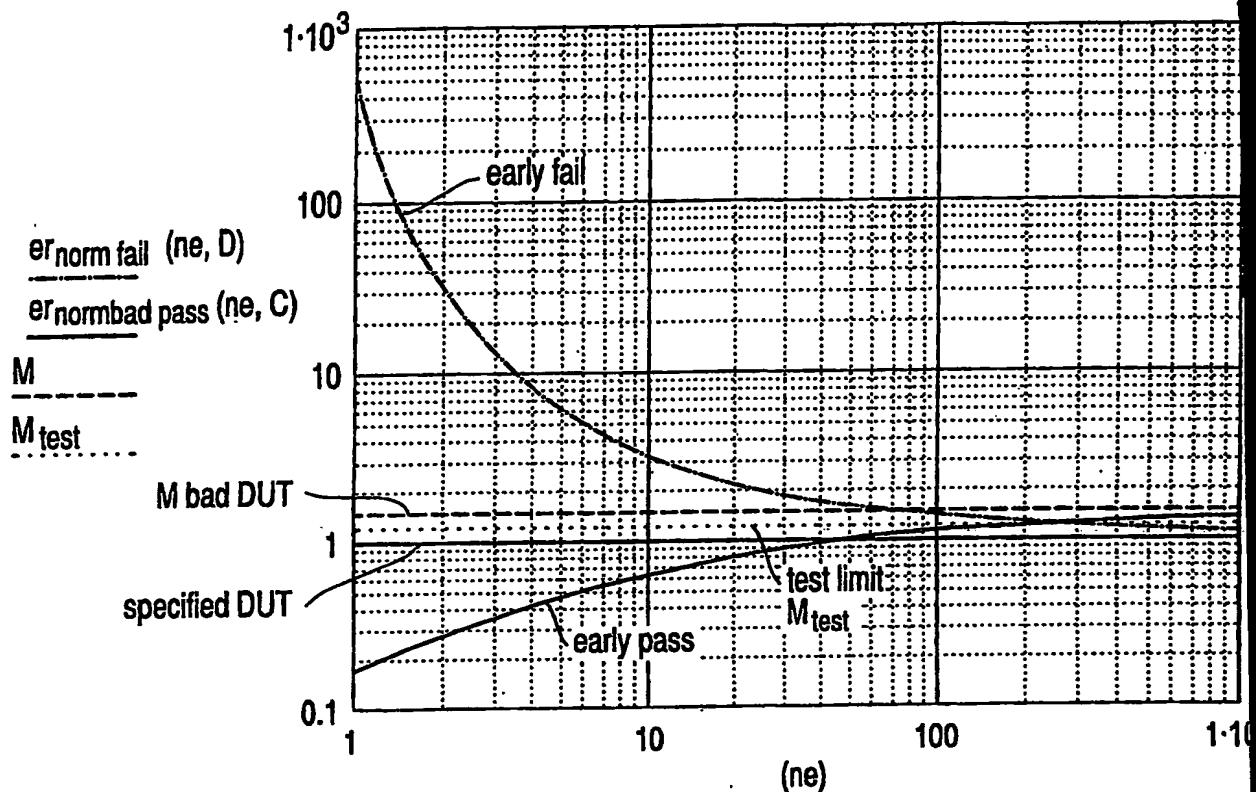


Fig. 6

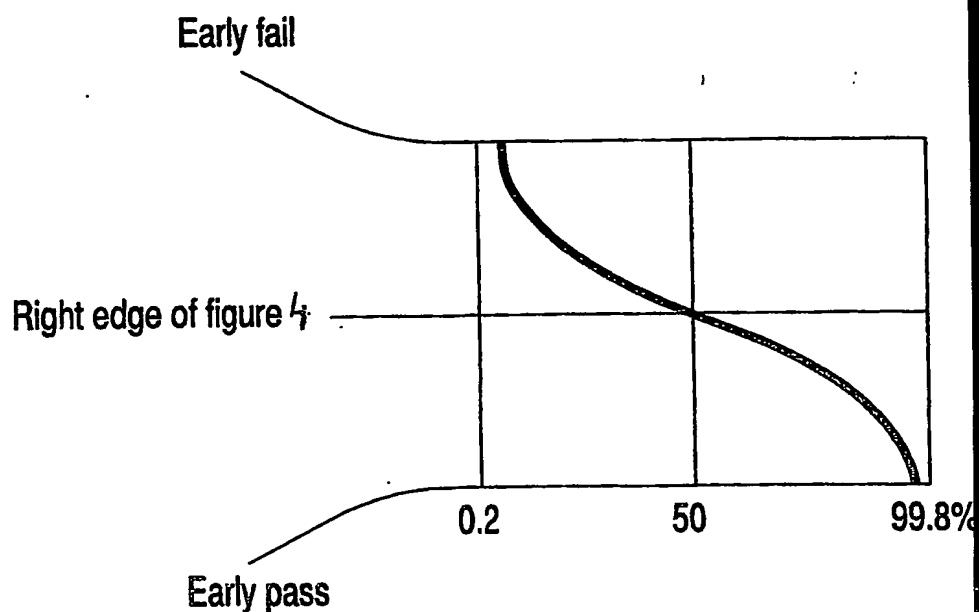
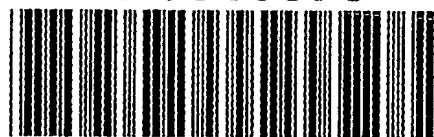


Fig. 7

PCT Application  
**EP0310873**



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